

# 9. Non-parametric statistics - Kruskal Wallis

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## 1 Comparison of $g$ groups

- Extend  $F$ -test from a one-way ANOVA to non-parametric alternatives.

## 2 DMH example

Assess genotoxicity of 1,2-dimethylhydrazine dihydrochloride (DMH) (EU directive)

- 24 rats
- four groups with daily DMH dose
  - control
  - low
  - medium
  - high
- Genotoxicity in liver using comet assay on 150 liver cells per rat.
- Are there differences in DNA damage due to DMH dose?

### 2.1 Comet Assay:

- Visualise DNA strand breaks
- Length comet tail is a proxy for strand breaks.

```
dna <- read_delim("https://raw.githubusercontent.com/GTPB/PSLS20/master/data/dna.txt", delim = " ")
dna$dose <- as.factor(dna$dose)
dna
```

```
# A tibble: 24 x 3
  id   length dose
<chr> <dbl> <fct>
1 Rat1  19.3  0
2 Rat2  18.9  0
3 Rat3  18.6  0
4 Rat4  19.0  0
```

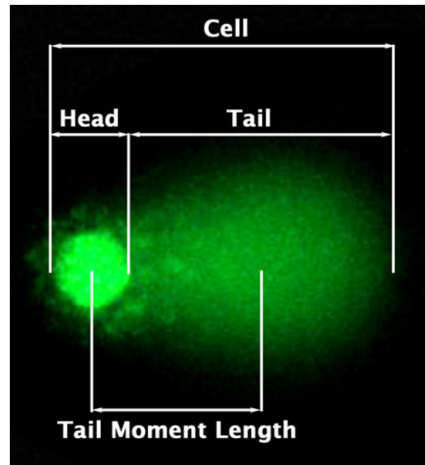
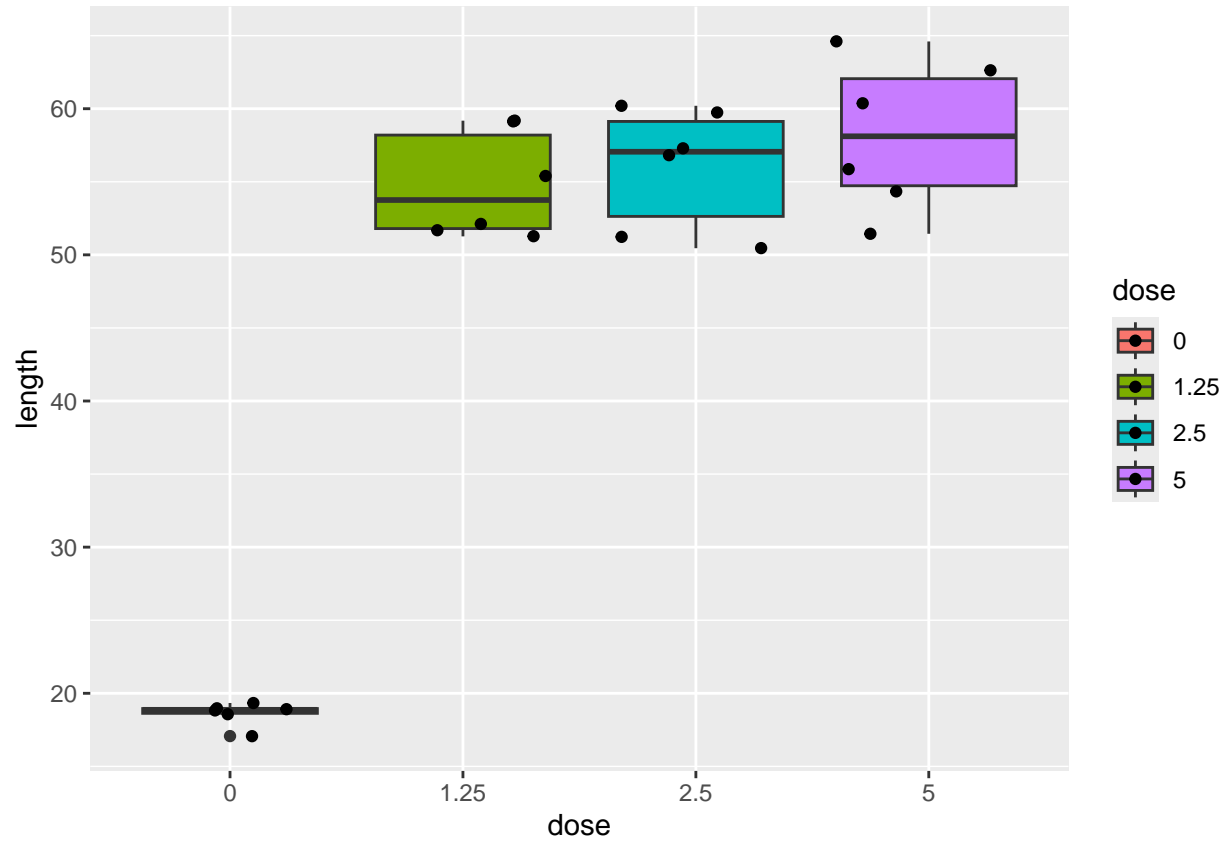


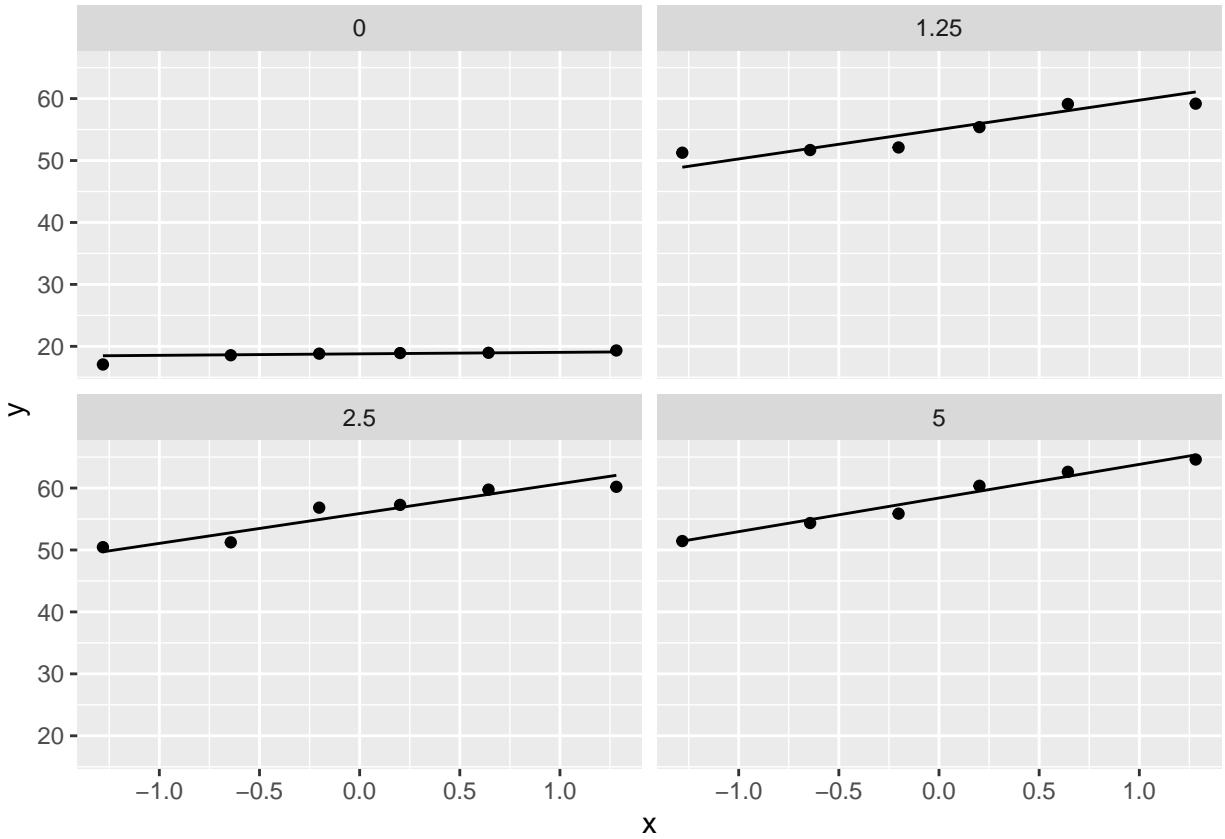
Figure 1: Comet assay

```
5 Rat5    17.1 0
6 Rat6    18.8 0
7 Rat7    55.4 1.25
8 Rat8    59.2 1.25
9 Rat9    59.1 1.25
10 Rat10  52.1 1.25
# i 14 more rows
```

```
dna %>%
  ggplot(aes(x = dose, y = length, fill = dose)) +
  geom_boxplot() +
  geom_point(position = "jitter")
```

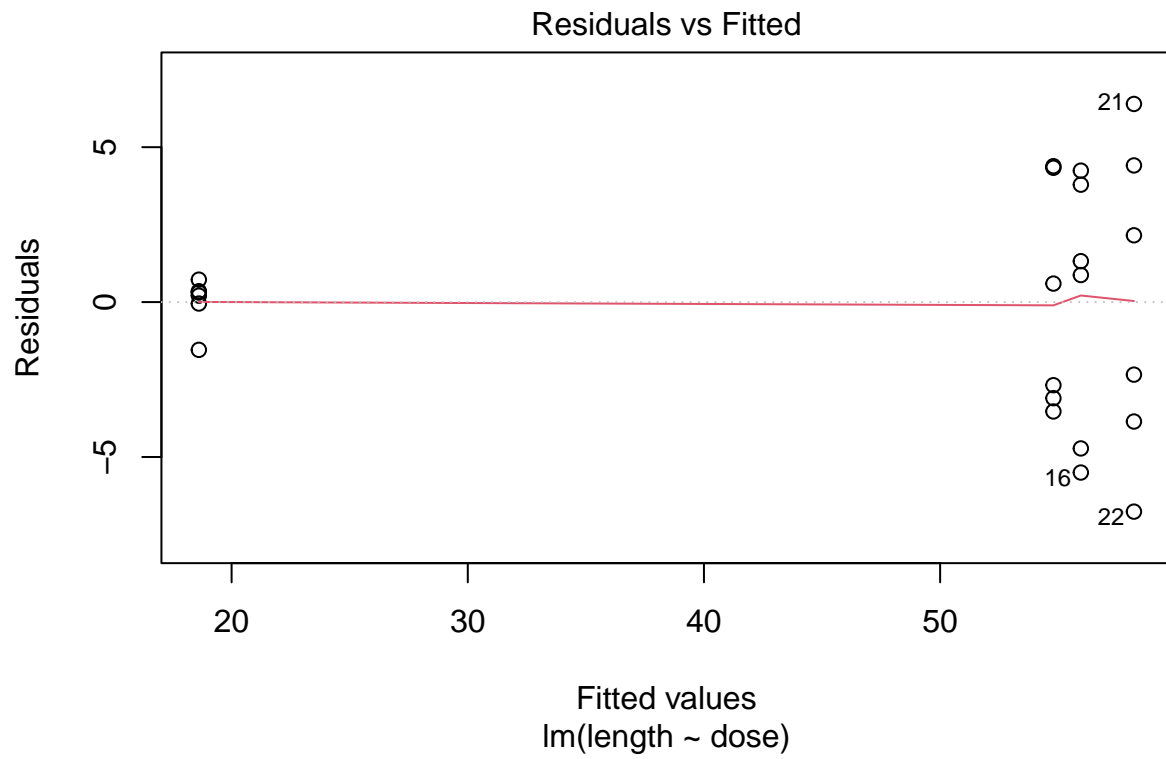


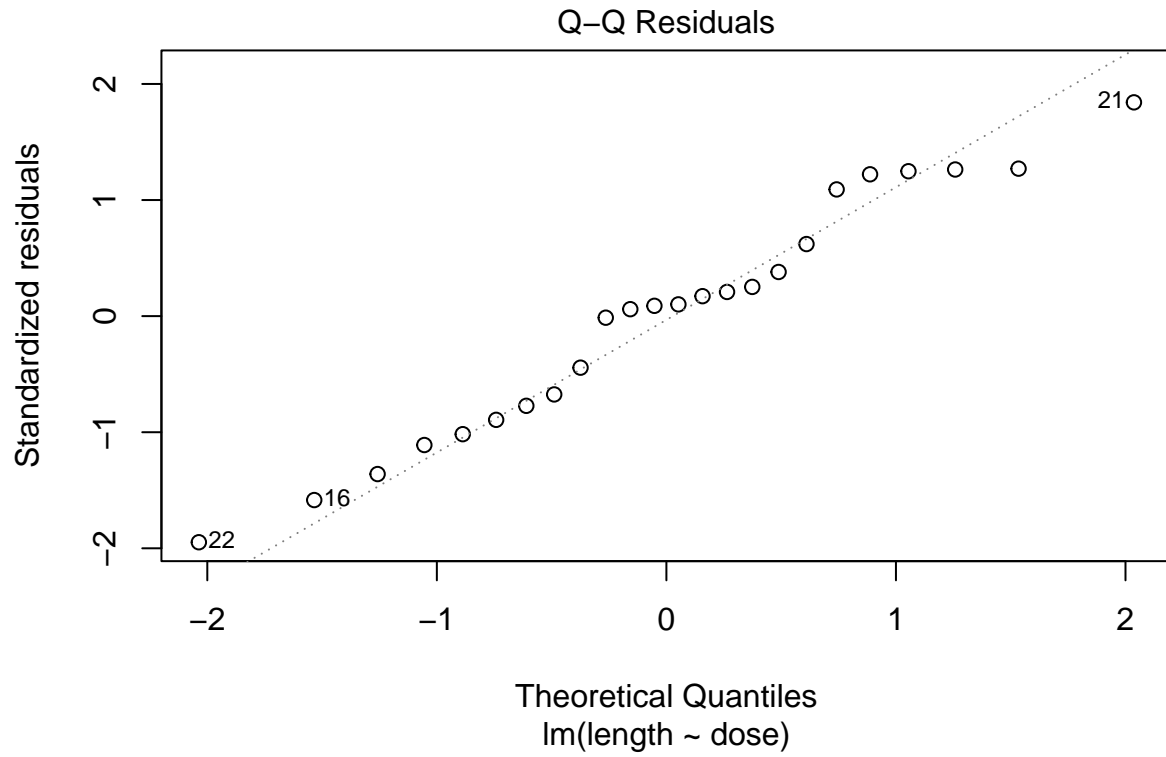
```
dna %>%
  ggplot(aes(sample = length)) +
  geom_qq() +
  geom_qq_line() +
  facet_wrap(~dose)
```

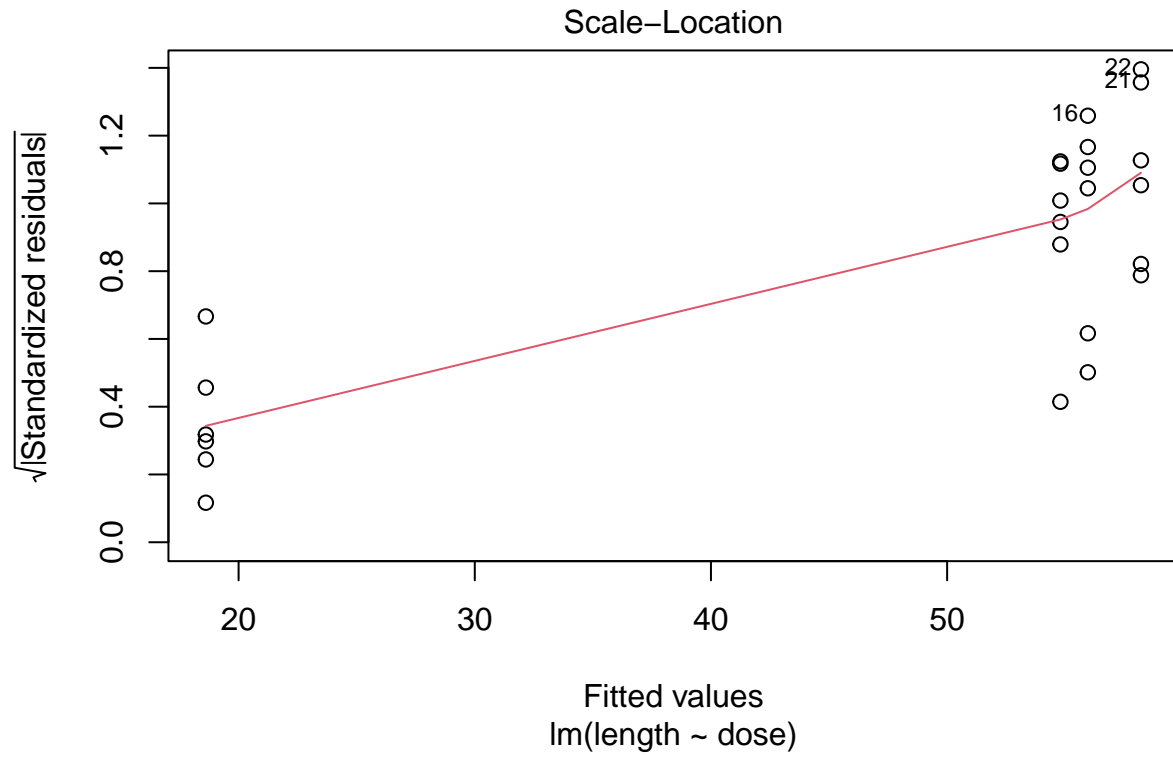


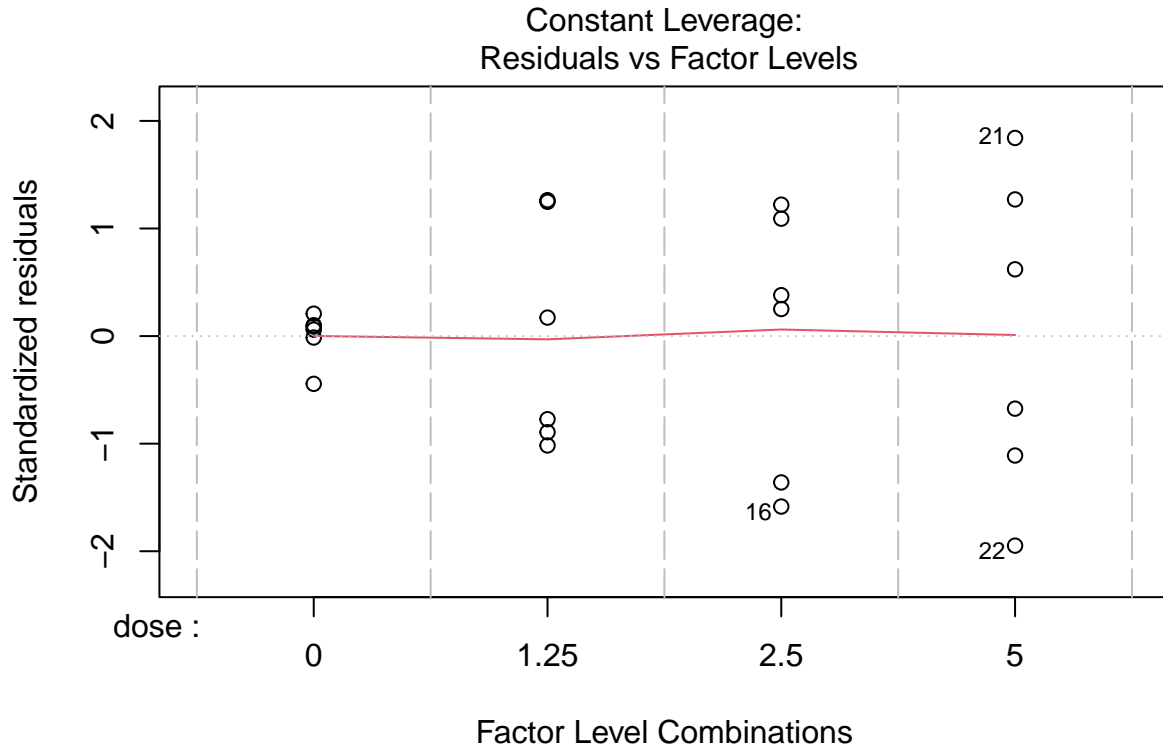
- Strong indication that data in control group has a lower variance.
- 6 observations per group are too few to check the assumptions

```
plot(lm(length ~ dose, data = dna))
```









### 3 Kruskal-Wallis Rank Test

- The Kruskal-Wallis Rank Test (KW-test) is a non-parametric alternative for ANOVA F-test.
- Classical  $F$ -test statistic can be written as

$$F = \frac{SST/(g-1)}{SSE/(n-g)} = \frac{SST/(g-1)}{(SSTot - SST)/(n-g)},$$

- with  $g$  the number of groups.
- $SSTot$  depends only on outcomes  $y$  and will not vary in permutation test.
- $SST$  can be used as statistic

$$SST = \sum_{j=1}^t n_j (\bar{Y}_j - \bar{Y})^2$$

- The KW test statistic is based on  $SST$  on rank-transformed outcomes<sup>1</sup>,

$$SST = \sum_{j=1}^g n_j (\bar{R}_j - \bar{R})^2 = \sum_{j=1}^t n_j \left( \bar{R}_j - \frac{n+1}{2} \right)^2,$$

- with  $\bar{R}_j$  the mean of the ranks in group  $j$ , and  $\bar{R}$  the mean of all ranks,

$$\bar{R} = \frac{1}{n}(1 + 2 + \dots + n) = \frac{1}{n} \frac{1}{2} n(n+1) = \frac{n+1}{2}.$$

<sup>1</sup>we assume that no *ties* are available



- The KW teststatistic is given by

$$KW = \frac{12}{n(n+1)} \sum_{j=1}^g n_j \left( \bar{R}_j - \frac{n+1}{2} \right)^2.$$

- The factor  $\frac{12}{n(n+1)}$  is used so that  $KW$  has a simple asymptotic null distribution. In particular under  $H_0$ , given thart  $\min(n_1, \dots, n_g) \rightarrow \infty$ ,

$$KW \rightarrow \chi_{t-1}^2.$$

- The exact KW-test can be executed by calculating the permutation null distribution (that only depends on  $n_1, \dots, n_g$ ) to test

$$H_0 : f_1 = \dots = f_g \text{ vs } H_1 : \text{ at least two means are different.}$$

- In order to allow  $H_1$  to be formulated in terms of means, the assumption of locations shift should be valid.
- For DMH example this is not the case.
- If location-shift is invalid, we have to formulate  $H_1$  in terms of probabilistic indices:

$$H_0 : f_1 = \dots = f_g \text{ vs } H_1 : \exists j, k \in \{1, \dots, g\} : P [Y_j \geq Y_k] \neq 0.5$$

### 3.1 DNA Damage Example

```
kruskal.test(length ~ dose, data = dna)
```

Kruskal-Wallis rank sum test

```
data: length by dose
Kruskal-Wallis chi-squared = 14, df = 3, p-value = 0.002905
```

- On the 5% level of significance we can reject the null hypothesis.
- R-functie `kruskal.test` only returns the asymptotic approximation for  $p$ -values.
- With only 6 observaties per groep, this is not a good approximation of the  $p$ -value
- With the `coin` R package we can calculate the exacte  $p$ -value

```
library(coin)
kwPerm <- kruskal_test(length ~ dose,
  data = dna,
  distribution = approximate(B = 100000)
)
kwPerm
```

Approximative Kruskal-Wallis Test

```
data: length by dose (0, 1.25, 2.5, 5)
chi-squared = 14, p-value = 0.00043
```

- We conclude that the difference in the distribution of the DNA damages due to the DMH dose is extremely significantly different.
- Posthoc analysis with WMW tests.

```
pairwise.wilcox.test(dna$length, dna$dose)
```

Pairwise comparisons using Wilcoxon rank sum exact test

data: dna\$length and dna\$dose

```
      0      1.25  2.5
1.25 0.013 -      -
2.5  0.013 0.818 -
5     0.013 0.721 0.788
```

P value adjustment method: holm

- All DMH behandelingen are significantly different from the control.
- The DMH are not significantly different from one another.
- U1 does not occur in the `pairwise.wilcox.test` output. Point estimate on probability on higher DNA-damage?

```
nGroup <- table(dna$dose)
probInd <- combn(levels(dna$dose), 2, function(x) {
  test <- wilcox.test(length ~ dose, subset(dna, dose %in% x))
  return(test$statistic / prod(nGroup[x]))
})
names(probInd) <- combn(levels(dna$dose), 2, paste, collapse = "vs")
probInd
```

```
      0vs1.25      0vs2.5      0vs5 1.25vs2.5  1.25vs5  2.5vs5
0.0000000 0.0000000 0.0000000 0.4444444 0.2777778 0.3333333
```

Because there are doubts on the location-shift model we draw our conclusions in terms of the probabilistic index.

### 3.1.1 Conclusion

- There is an extremely significant difference in in the distribution of the DNA-damage measurements due to the treatment with DMH ( $p < 0.001$  KW-test).
- DNA-damage is more likely upon DMH treatment than in the control treatment (all  $p=0.013$ , WMW-testen).
- The probability on higher DNA-damage upon exposure to DMH is 100% (Calculation of a CI on the probabilistic index is beyond the scope of the course)
- There are no significant differences in the distributions of the comit-lengths among the treatment with the different DMH concentrations ( $p = 0.72-0.82$ ).
- DMH shows already genotoxic effects at low dose.
- (Alle paarswise tests are gecorrected for multiple testing using Holm's methode).